

## APMA 1650 – MIDTERM 2 – SOLUTIONS

1. (a)

$$\begin{aligned}\int_{-\infty}^{\infty} f(y) dy &= \int_0^2 c(2-y) dy \\ &= \int_0^2 2c - cy dy \\ &= \left[ 2cy - c \left( \frac{y^2}{2} \right) \right]_0^2 \\ &= 2c(2) - c \left( \frac{2^2}{2} \right) \\ &= 4c - 2c \\ &= 2c = 1\end{aligned}$$

$$\boxed{c = 1/2}$$

(b)

$$\begin{aligned}E(Y) &= \int_{-\infty}^{\infty} yf(y) dy \\ &= \int_0^2 y \left[ \frac{1}{2}(2-y) \right] dy \\ &= \int_0^2 y - \frac{1}{2}y^2 dy \\ &= \left[ \frac{y^2}{2} - \frac{1}{6}y^3 \right]_0^2 \\ &= \frac{4}{2} - \frac{8}{6} \\ &= 2 - \frac{4}{3} \\ &= \frac{2}{3}\end{aligned}$$

$$\boxed{E(Y) = \frac{2}{3}}$$

2. (a) Let  $Y$  be the student's score.

Since we know  $E(Y) = 50$  by Markov's inequality with  $a = 65$ , we have

$$P(Y \geq 65) \leq \frac{E(Y)}{65} = \frac{50}{65}$$

$$P(Y < 65) = 1 - P(Y \geq 65) = 1 - \frac{50}{65} = \frac{15}{65} = \frac{3}{13}$$

**Answer:**  $\boxed{3/13}$

- (b) Suppose  $\text{Var}(Y) = \sigma^2$

Then by Chebyshev's inequality with  $a = 10$  we have

$$P(|Y - E(Y)| \geq 10) \leq \frac{\text{Var}(Y)}{10^2}$$

$$P(|Y - 50| \geq 10) \leq \frac{\sigma^2}{100}$$

$$\begin{aligned} P(40 \leq Y \leq 60) &= P(-10 \leq Y - 50 \leq 10) \\ &= P(|Y - 50| \leq 10) \\ &= 1 - P(|Y - 50| \geq 10) \\ &\geq 1 - \frac{\sigma^2}{100} \end{aligned}$$

In order to ensure that this is  $\geq 0.64$  it's enough to have

$$\begin{aligned} 1 - \frac{\sigma^2}{100} &\geq 0.64 \\ -\frac{\sigma^2}{100} &\geq -0.36 \\ \frac{\sigma^2}{100} &\leq 0.36 \\ \sigma^2 &\leq 36 \\ \sigma &\leq 6 \end{aligned}$$

Therefore the largest standard deviation allowed is  $\boxed{\sigma = 6}$

3. (a)

$$\begin{aligned} E[\hat{\theta}_3] &= E[a\hat{\theta}_1 + (1-a)\hat{\theta}_2] \\ &= aE[\hat{\theta}_1] + (1-a)E[\hat{\theta}_2] \\ &= a\theta + (1-a)\theta \\ &= \theta \end{aligned}$$

Hence  $\hat{\theta}_3$  is an unbiased estimator of  $\theta$

(b) By the Magic MSE formula, we have

$$\text{MSE}(\hat{\theta}_3) = [\text{Bias}(\hat{\theta}_3)]^2 + \text{Var}(\hat{\theta}_3)$$

Since  $\hat{\theta}_3$  is unbiased we have  $\text{Bias}(\hat{\theta}_3) = 0$  hence we just need to find  $\text{Var}(\hat{\theta}_3)$ .

But since  $\hat{\theta}_1$  and  $\hat{\theta}_2$  are independent, we have

$$\begin{aligned} \text{Var}(\hat{\theta}_3) &= \text{Var}(a\hat{\theta}_1 + (1-a)\hat{\theta}_2) \\ &= \text{Var}(a\hat{\theta}_1) + \text{Var}(1-a)\hat{\theta}_2 \\ &= a^2 \text{Var}(\hat{\theta}_1) + (1-a)^2 \text{Var}(\hat{\theta}_2) \\ &= a^2\sigma_1^2 + (1-a)^2\sigma_2^2 \end{aligned}$$

$$\text{MSE}(\hat{\theta}_3) = a^2\sigma_1^2 + (1-a)^2\sigma_2^2$$

4. Since the sample size is large, by the Central Limit Theorem,  $\bar{Y}$  is approximately normal.

Since we don't know  $\sigma$  we use  $\hat{\sigma} = \frac{S}{\sqrt{n}} = \frac{3}{10} = 0.3$

$$1 - \alpha = 0.6 \Rightarrow \alpha = 0.4 \Rightarrow \frac{\alpha}{2} = 0.2$$

And from the  $z$ -table we get  $-z_{\alpha/2} = -0.8 \Rightarrow z_{\alpha/2} = 0.8$  and therefore our confidence interval is

$$\begin{aligned} [\hat{L}, \hat{U}] &= [\bar{Y} - (z_{\alpha/2}) \hat{\sigma}, \bar{Y} + (z_{\alpha/2}) \hat{\sigma}] \\ &= [50 - (0.8)(0.3), 50 + (0.8)(0.3)] \\ &= [50 - 0.24, 50 + 0.24] \\ &= [49.76, 50.24] \end{aligned}$$

5. (a)

$$\begin{aligned}\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy &= \int_0^1 \int_{x^2}^1 cx dy dx \\ &= \int_0^1 cx (1 - x^2) dx \\ &= \int_0^1 cx - cx^3 dx \\ &= \left[ c \left( \frac{x^2}{2} \right) - c \left( \frac{x^4}{4} \right) \right]_0^1 \\ &= \frac{c}{2} - \frac{c}{4} \\ &= \frac{c}{4} = 1\end{aligned}$$

Hence  $\boxed{c = 4}$ 

(b)

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy = \int_{x^2}^1 4x dy = 4x (1 - x^2) = 4x - 4x^3$$

Since  $0 \leq x \leq 1$  we have

$$f_X(x) = \begin{cases} 4x - 4x^3 & \text{if } 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$